Data Analysis

The explained variance and informational redundancy

• The quantity of variance explained by each pair of canonical variables, in connection with each of the initial data set, is given by the sum of correlations between the canonical variables and the causal variables of the sets:

$$VX_{k} = \sum_{j=1}^{p} R(z_{k}, X_{j})^{2}, k = 1, m$$
$$VY_{k} = \sum_{j=1}^{q} R(u_{k}, Y_{j})^{2}, k = 1, m,$$

• where $R(z_k, X_j)^2$ is the determination (correlation) coefficient between the canonical variable z_k of the pair k, and the variable X_j , belonging to the first data set (the *j* column of matrix X),

The explained variance and informational redundancy

- and $R(u_k, Y_j)^2$ is the correlation coefficient between the canonical variable u_k of the pair k, and the causal variable Y_j , belonging to the second data set.
- Proportionally, the values are: $\frac{VX_k}{p}$ and $\frac{VY_k}{q}$, p and q

being the number of causal variables, columns, of matrices *X*, and *Y*, respectively.

• The overall variance explained by all *m* canonical roots is: $VX = \sum_{k=1}^{m} VX_k$, for the first data set *X*, and $VY = \sum_{k=1}^{m} VY_k$, for the second data set *Y*.

The explained variance and informational redundancy

- The redundancy is given by the common information existent in both data sets, and extracted by the canonical roots (pairs).
- The common information is given by the canonical correlation.
- If there is a certain quantity of information extracted by a canonical variable from one of the sets, then the part of this information found in the other set it is retrieved by using the canonical correlation, as follows:

$$SX_k = VX_k \cdot \alpha_k, \qquad k = 1, m,$$

$$SY_k = VY_k \cdot \alpha_k, \qquad k = 1, m,$$

where the eigenvalue α_k is the correlation coefficient between the canonical variables z_k and u_k .

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The explained variance and informational redundancy

• The redundancy of all *m* canonical roots is:

$$SX = \sum_{k=1}^{m} SX_{k},$$
$$SY = \sum_{k=1}^{m} SY_{k}.$$

Standardizing canonical factors

• Canonical factors are easier to interpret if standardized. Standardizing canonical factors implies to relate them to the standard deviation of initial and canonical variables.

$$as_{ik} = a_{ik} \cdot \frac{\sigma_{X_i}}{\sigma_{x_i}} , i = 1, p, k = 1, m$$
$$bs_{ik} = b_{ik} \cdot \frac{\sigma_{Y_i}}{\sigma_{x_i}} , i = 1, q, k = 1, m$$

• The interpretation of standardized factors is similar to multiple regression: the increase with one unit of variables X_i or Y_i standard deviation, generates an increase with as_{ik} or bs_{ik} of canonical variables z_k or u_k standard deviation.

Canonical roots - Bartlett χ^2 relevance test

- Bartlett χ^2 is the most employed test to evaluate canonical correlations.
- For any given canonical root, the result of the test indicates if there is any dependency between the two sets of variables or, on the contrary, the two sets of variables are independent.
- H0 hypothesis: correlation coefficient $R(z_k,u_k)$ indicates the existence of a linear correlation between the two sets of initial (causal) variables.
- The alternative hypothesis, H1: correlation coefficient $R(z_k, u_k)$ indicates a lack of connection.

Canonical roots - Bartlett χ^2 relevance test

• For a canonical root (z_k, u_k) , the test is applied as follows:

1. The number of degrees of freedom is computed, associated to each canonical root of rank *k*:

$$df_k = (p - k + 1)(q - k + 1)$$

where p and q are the number of initial variables of the first and second sets, respectively.

2. The statistics of the test is computed as follows:

$$\chi_k^2 = \left(-n+1+\frac{p+q+1}{2}\right)\log(1-\lambda_k)$$

where *n* is the number of observations, and λ_k is an indicator named lambda Wilks.

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Canonical roots - Bartlett χ^2 relevance test

Lambda Wilks indicator is computed in the following manner:

 $\lambda_k = \prod_{i=k}^m \left(1 - R(z_i, u_i)^2 \right) \quad \text{, where } m \text{ is the number of canonical}$

roots.

3. Using χ^2 distribution, it is then determined the critical value for the test:

 $\chi c_k^2 (1 - \alpha, df_k)$, for a significance threshold, α .

4. Then the test is applied:

If $\chi_k^2 \geq \chi c_k^2 (1-\alpha, df_k)$,

the H0 hypothesis is accepted, with a level of credence $1-\alpha$, otherwise it is rejected.

Generalized Canonical Analysis (gCCA)

Having given q data sets $X_1, X_2, ..., X_q$ which describe the same n observations, m_i the number of columns of matrix X_i and W_i the subspace in \mathbb{R}^n generated by the columns, we make the following notations:

- P_i is the orthogonal projection X_i on subspace W_i .
- The total number of causal variables is $m = \sum_{i=1}^{7} m_i$.
- Make the assumption that n > m.

Generalized Canonical Analysis (gCCA)

Generalized canonical analysis is to determine in the first phase an auxiliary variable Z_1 , as a linear combination of causal variables and q canonical variables z_{i1} (i = 1, q), such that:

- $\sum_{i=1}^{q} R^2(Z_1, z_{i1})$ to be maximal, under the restriction of having:
- $(Z_1)^t Z_1 = 1$

Generalized Canonical Analysis (gCCA)

- In order to have the sum of the correlation maximal, the vectors z_{i1} are selected such that to be orthogonal projection of Z_1 vector on subspaces W_i : $z_{i1} = P_i \cdot Z_1$.
- Returning to the correlation sums, it can be rewritten as follows:

$$\sum_{i=1}^{q} R^{2}(Z_{1}, z_{i1}) = \sum_{i=1}^{q} \frac{Cov(Z_{1}, z_{i1})^{2}}{Var(Z_{1})Var(z_{i1})} =$$

$$= \sum_{i=1}^{q} \frac{\left(\frac{1}{n}(Z_{1})^{t} z_{i1}\right)^{2}}{\frac{1}{n}(Z_{1})^{t} Z_{1}} \frac{1}{n}(z_{i1})^{t} z_{i1}} = \sum_{i=1}^{q} \frac{\left(\frac{1}{n}(Z_{1})^{t} z_{i1}\right)^{2}}{\frac{1}{n^{2}}(z_{i1})^{t} z_{i1}} = \sum_{i=1}^{q} \frac{\left((Z_{1})^{t} z_{i1}\right)^{2}}{(Z_{i1})^{t} z_{i1}}$$

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Generalized Canonical Analysis (gCCA)

• Replacing z_{i1} with $P_i \cdot Z_1$, we obtain:

$$\sum_{i=1}^{q} R^{2}(Z_{1}, z_{i1}) = \sum_{i=1}^{q} \frac{\left(\left(Z_{1} \right)^{t} P_{i} Z_{1} \right)^{2}}{\left(Z_{1} \right)^{t} \left(P_{i} \right)^{t} P_{i} Z_{1}} = \sum_{i=1}^{q} \left(Z_{1} \right)^{t} P_{i} Z_{1} = \left(Z_{1} \right)^{t} \left(\sum_{i=1}^{q} P_{i} \right) Z_{1},$$

because
$$(P_i)^t P_i = P_i^2 = P_i$$

• Therefore, the optimum problems becomes:

$$\begin{cases} Maxim(Z_1)^t \left(\sum_{i=1}^q P_i\right) Z_1 \\ (Z_1)^t Z_1 = 1 \end{cases}$$

Generalized Canonical Analysis (gCCA)

• The solution of this problem (see PCA), the variable Z_1 , is the eigenvector of the matrix $\sum_{i=1}^{q} P_i$

corresponding to the greatest eigenvalue, while the canonical variables of the sets, z_{i1} , are determined using the relation:

 $z_{i1} = P_i \cdot Z_1.$

• At the *k* phase is to be determined the auxiliary variable Z_k and the canonical variables z_{ik} (*i* = 1, *q*) such that: $\sum_{k=1}^{q} R^2(Z_k, z_{ik})$

to be maximal.

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Generalized Canonical Analysis (gCCA)

• Under the following conditions:

1)
$$(Z_k)^t Z_k = 1$$
,
2) $(Z_k)^t Z_j = 0$, $j = \overline{1, k - 1}$.

• The variable Z_k is the eigenvector of matrix $\sum_{i=1}^{4} P_i$,

corresponding to the eigenvalue of rank k, while the canonical variables of the sets are: $z_{ik} = P_i Z_k$.

Generalized Canonical Analysis (gCCA)

• The orthogonal projection on space W_i is determined using the following relation:

 $P_i = X_i (X_i^{t} X_i)^{-1} X_i^{t}, \quad i=1,q$ (see the canonical analysis).

• If the sum of the projection matrices is developed then: $\sum_{i=1}^{q} P_i = \sum_{i=1}^{q} X_i \cdot (X_i^{t} X_i)^{-1} X_i^{t} = X \cdot D_{XX}^{-1} X^{t},$

•

where D_{XX} is a block-diagonal matrix of the following format:

$$\begin{bmatrix} V_{11} & 0 & \dots & 0 \\ 0 & V_{22} & \dots & 0 \\ \dots & & \dots & \\ 0 & 0 & \dots & V_{qq} \end{bmatrix}$$

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Generalized Canonical Analysis (gCCA)

with $V_{jj} = X_j^t X$.

- This matrix $X \cdot D_{XX}^{-1} X^{t}$ has *n* rows and *n* columns.
- The number non-zero eigenvalue of the matrix $X \cdot D_{XX}^{-1} X^{t}$, and the number of implicit steps (phases) of the algorithm is *m*.
- An auxiliary variable Z_k is eigenvector of the matrix $X \cdot D_{XX}^{-1} X^t$ if :

$$X \cdot D_{XX}^{-1} X^{t} Z_{k} = \alpha_{k} Z_{k}$$
⁽¹⁾

Generalized Canonical Analysis (gCCA)

• Since Z_k is a linear combination of causal variables, it can be written as:

$$Z_k = X \cdot b_k \quad ,$$

where b_k is the vectorul of the linear combination or *the factor*.

- Replacing Z_k in relation (1) it is obtained: $X \cdot D_{XX}^{-1} X^{t} X \cdot b_k = \alpha_k X \cdot b_k$ (2)
- Multiplying relation (2) at the left with the matrix $(X^{t}X)^{-1}X^{t}$ it is obtained:

$$\left(X^{t}X\right)^{-1}X^{t}X\cdot D_{XX}^{-1}X^{t}X\cdot b_{k} = \alpha_{k}\left(X^{t}X\right)^{-1}X^{t}X\cdot b_{k}$$

Generalized Canonical Analysis (gCCA)

- Therefore: $D_{XX}^{-1} X^{t} X \cdot b_{k} = \alpha_{k} b_{k}$.
- Hence the factors b_k are obtained as eigenvectors of the matrix $D_{XX}^{-1} X^{t} X$
- The eigenvalues of the matrix $D_{XX}^{-1} X^{t} X$ coincide with the *m* non-zero eigenvalues of the matrix $X \cdot D_{XX}^{-1} X^{t}$.
- The eigenvalues represent the sum of the determination (correlation) coefficients between the auxiliary variables Z_k and the canonical variables of the sets:

$$\alpha_k = \sum_{j=1}^q R(Z_k, z_{jk})^2 \quad , \quad k = \overline{1, m} \, .$$

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Generalized Canonical Analysis (gCCA) - *numerical example*

- Having given the following observation tables, the number of causal variable sets is q=3.
- The sets are $(X^1, X^2, X^3, X^4, X^5)$, (Y^1, Y^2, Y^3, Y^4) , and (V^1, V^2, V^3) .

Generalized Canonical Analysis (gCCA) - *numerical example*

X1	X2	Х3	X4	X5	Y1	Y2	Y3	Y4	V1	V2	V3
2	12	1.5	9	16	1.5	19.5	2	27.5	1	0	0
12	7	2.5	19	32	7.5	9	1	8	1	0.5	0.5
16	1	2.5	14	42	7.5	9	0	5.5	1	0.5	0
5.5	17.5	1	51	5.5	1.5	2.5	1	4	0	0	10
12	3	4.5	4	61	4	4	1	5	1	0.5	0
3	55.5	0	15	3.5	1.5	4	0.5	12	1	2	2
4	56	0	16	3	2	3.5	0.5	10	0.5	0.5	0.5
0	71	2	12	2	1	1	0	9	0	0	0
2.5	31	2.5	13.5	6	1	8.5	0	22.5	2	2	2.5
4	20	2	17	11	5	11	3	23.5	1.5	0.5	0.5
5.5	37	1	24	8.5	2.5	6	1	10	0.5	1	0.5
11	9	4	34	22	5	6	0	6	1	1.5	0.5
0	8	0	3	9	9	37	6	20	5	0	0
34	2	5.5	15	33	2.5	3	0	2.5	0	1	1
14.5	0	2	9	21.5	0	2	0	0.5	0	49.5	0
14	2.5	4.5	4.5	67	1.5	3.5	0	2	0.5	0.5	0
15	3	0.5	32	8	0	1	0	2.5	3.5	0.5	34
28	0	0.5	5	3.5	0	0	0	0.5	0	0	61.5
8.5	3	3.5	3	58.5	5.5	6	1.5	7.5	1	0.5	0

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Generalized Canonical Analysis (gCCA) - *numerical example*

• The following table presents the determining ratios between the synthetic variables, the canonical variables of the sets and the their sums (the eigenvalues).

Generalized Canonical Analysis (gCCA) - *numerical example*

	Set 1	Set 2	Set 3	Total (α_k)
Root 1	0.83593	0.88593	0.77377	2.49562
Root 2	0.82947	0.62577	0.88255	2.33779
Root 3	0.78863	0.82285	0.04012	1.65161
Root 4	0.55578	0.46533	0.35026	1.37137
Root 5	0.58676	0.16769	0.29233	1.04678
Root 6	0.55483	0.12036	0.30009	0.97528
Root 7	0.36705	0.36119	0.02172	0.74996
Root 8	0.19657	0.21633	0.06332	0.47622
Root 9	0.16745	0.22798	0.06975	0.46518
Root 10	0.04156	0.07389	0.17249	0.28793
Root 11	0.07568	0.0324	0.03337	0.14145
Root 12	0.00029	0.00028	0.00024	0.00082

Generalized Canonical Analysis (gCCA) - *numerical example*

- The correlation circle between the synthetic variables and set variables, highlights the homogeneity of the sets.
- The following graphic depicts the correlation circle between the first two canonical roots (Z_1, Z_2) .

Generalized Canonical Analysis (gCCA) - numerical example



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Generalized Canonical Analysis (gCCA) - *numerical example*

- The following graphic shows the correlation circle between the canonical variables of the sets and the initial, causal variables.
- This is an output with values in all 3 spaces.
- A grouped, compact distribution on the graphic for each variable, indicates the homogeneity of the sets.

Generalized Canonical Analysis (gCCA) - *numerical example*

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